

**Q.1** (20 points) Let  $X$  and  $Y$  be random variables. The joint density function of  $X$  and  $Y$  is given by

$$f(x, y) = \begin{cases} 8xy & \text{if } 0 \leq x \leq y \leq 1; \\ 0 & \text{otherwise.} \end{cases}$$

(a) Find the marginal density  $f_X(x)$  of  $X$ .

$$f_X(x) = \int_x^1 8xy \, dy = 4x(1 - x^2), \quad 0 \leq x \leq 1.$$

(b) Find the conditional density  $f(y|x)$  of  $Y$  given  $X = x$ .

$$f(y|x) = \begin{cases} \frac{2y}{1-x^2} & \text{if } x \leq y \leq 1; \\ 0 & \text{otherwise.} \end{cases}$$

(c) Find the best predictor of  $Y$  given  $X$ .

$$E(Y|X = x) = \int_{-\infty}^{\infty} yf(y|x)dy = \int_x^1 y \left( \frac{2y}{1-x^2} \right) dy = \frac{2(1-x^3)}{3(1-x^2)} = \frac{2(1+x+x^2)}{3(1+x)}$$

$$\text{Thus, } E(Y|X) = \frac{2(1-X^3)}{3(1-X^2)} = \frac{2(1+X+X^2)}{3(1+X)}.$$

**Q.2** (20 points) The function  $f(x, y) = \frac{1}{16\pi} \exp(-\frac{1}{2}Q(x, y))$  with the quadratic form

$$Q(x, y) = (x - 7)^2 + \left( \frac{y - 6x - 58}{8} \right)^2$$

gives the joint density function for  $(X, Y)$ .

(a) Find the marginal density function  $f_X(x)$  for  $X$ .

$$f_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{1}{2}(x-7)^2\right]$$

(b) Find the conditional density function  $f_{Y|X}(y|x)$ .

$$f_{Y|X}(y|x) = \frac{1}{8\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left(\frac{y-6x-58}{8}\right)^2\right\}$$

(c) Find the conditional expectation  $E(Y|X)$  and the conditional variance  $\text{Var}(Y|X)$ .

$$E(Y|X) = 6X + 58 \text{ and } \text{Var}(Y|X) = (8)^2$$

**Q.3** (20 points) Suppose that  $X_1$  and  $X_2$  are independent normal random variables with parameters  $(\mu_1, \sigma_1^2)$ ,  $(\mu_2, \sigma_2^2)$ , and that  $\mu_1 = 3$  and  $\mu_2 = 1$ . Consider the linear transformation  $U = g_1(X_1, X_2) = X_1 + X_2 + 5$  and  $V = g_2(X_1, X_2) = 2X_2 + 9$ . Then we know that  $(U, V)$  has a bivariate normal distribution with parameters  $\mu_u, \mu_v, \sigma_u, \sigma_v$ , and  $\rho$ .

- (a) We can express  $\begin{bmatrix} U \\ V \end{bmatrix} = A \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} + \mathbf{b}$ . Find the matrix  $A$  and the vector  $\mathbf{b}$ .

$$A = \begin{bmatrix} 1 & 1 \\ 0 & 2 \end{bmatrix} \text{ and } \mathbf{b} = \begin{bmatrix} 5 \\ 9 \end{bmatrix}$$

- (b) Find the mean parameters  $\mu_u$  and  $\mu_v$ .

$$\begin{bmatrix} \mu_u \\ \mu_v \end{bmatrix} = \begin{bmatrix} 1 & 1 \\ 0 & 2 \end{bmatrix} \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix} + \mathbf{b} = \begin{bmatrix} 1 & 1 \\ 0 & 2 \end{bmatrix} \begin{bmatrix} 3 \\ 1 \end{bmatrix} + \begin{bmatrix} 5 \\ 9 \end{bmatrix} = \begin{bmatrix} 9 \\ 11 \end{bmatrix}$$

- (c) Find  $\sigma_u$ ,  $\sigma_v$ , and  $\rho$  in terms of  $\sigma_1$  and  $\sigma_2$ .

The covariance matrix for  $(U, V)$  is given by

$$A\Sigma A^T = \begin{bmatrix} 1 & 1 \\ 0 & 2 \end{bmatrix} \begin{bmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 1 & 2 \end{bmatrix} = \begin{bmatrix} \sigma_1^2 + \sigma_2^2 & 2\sigma_2^2 \\ 2\sigma_2^2 & 4\sigma_2^2 \end{bmatrix}$$

Thus,  $\sigma_u = \sqrt{\sigma_1^2 + \sigma_2^2}$ ,  $\sigma_v = 2\sigma_2$  and  $\rho = \frac{\sigma_2}{\sqrt{\sigma_1^2 + \sigma_2^2}}$ .

*Remark.* The direct calculation of  $\text{Cov}(U, V)$  can be obtained as follows:

$$\begin{aligned} \text{Cov}(U, V) &= E[(U - \mu_u)(V - \mu_v)] = E[(X_1 + X_2 - 4)(2X_2 - 2)] \\ &= 2E[(X_1 - \mu_1)(X_2 - \mu_2)] + 2E[(X_2 - \mu_2)^2] = 2\text{Var}(X_2) = 2\sigma_2^2 \end{aligned}$$

$\text{Var}(X)$  and  $\text{Var}(Y)$  can be similarly calculated.

**Q.4** (20 points) Let

$$f(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1}(1-x)^{\beta-1}, \quad 0 < x < 1,$$

be a beta distribution with parameter  $(\alpha, \beta)$ . Answer the following questions.

- (a) Suppose that  $X_1, \dots, X_n$  are iid random variables from the uniform distribution having the pdf  $f(x) = 1$ ,  $0 < x < 1$ , and that  $X_{(k)}$  is the  $k$ -th order statistic. Find  $X_{(k)}$  having a beta distribution, and identify  $\alpha$  and  $\beta$  in terms of  $n$  and  $k$ .

The  $k$ -th order statistic  $X_{(k)}$  has the pdf  $f(x) = \frac{n!}{(k-1)!(n-k)!} x^{k-1}(1-x)^{n-k}$ . Thus,  $\alpha = k$  and  $\beta = n - k + 1$ .

- (b) Find  $E[X_{(k)}^3]$  in terms of  $n$  and  $k$ .

$$\begin{aligned} E[X_{(k)}^3] &= \int_0^1 x^3 f(x) dx = \frac{\Gamma(n+1)}{\Gamma(k)\Gamma(n-k+1)} \int_0^1 x^{k+2}(1-x)^{n-k} dx \\ &= \frac{\Gamma(n+1)}{\Gamma(k)\Gamma(n-k+1)} \left( \frac{\Gamma(k+3)\Gamma(n-k+1)}{\Gamma(n+4)} \right) = \frac{(k+2)(k+1)k}{(n+3)(n+2)(n+1)} \end{aligned}$$

**Q.5** (20 points) Let  $X_1, \dots, X_n$  be iid normal random variables with mean  $\mu$  and variance  $\sigma^2$ , and let  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$  and  $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$ .

- (a)  $k \times \left(\frac{S^2}{\sigma^2}\right)$  is related to some type of distribution by multiplying it by some constant value  $k$ .  
Give the name of distribution with specific degrees of freedom, and find  $k$  in terms of  $n$ .

It has a  $\chi^2$ -distribution with  $n - 1$  degrees of freedom when  $k = n - 1$ .

- (b)  $k \times \left(\frac{\bar{X} - \mu}{S}\right)$  is related to some type of distribution by multiplying it by some constant value  $k$ .  
Give the name of distribution with specific degrees of freedom, and find  $k$  in terms of  $n$ .

It has a  $t$ -distribution with  $n - 1$  degrees of freedom when  $k = \sqrt{n}$ .

- (c) Suppose that  $n = 9$ . Then find the value  $k$  so that  $P\left(\frac{\bar{X} - \mu}{S} < k\right) = 0.05$ .

$$k = -\frac{t_{0.05,8}}{\sqrt{9}} = -1.86/3 = -0.62.$$

- (d) Suppose that  $n = 25$ . Then find the value  $k$  so that  $P\left(\left|\frac{\bar{X} - \mu}{S}\right| \leq k\right) = 0.95$ .

$$k = \frac{t_{0.025,24}}{\sqrt{25}} = 2.064/5 = 0.4128.$$